## ECOM-G314 Econometrics 1 Homework Assignment 3

This version of the assignment includes some additional links to the course material to help you find where a particular topic is being discussed. It should otherwise be identical to the original assignment.

1. Consider the wage equation in the empirical example in the video lecture on IV estimator. The data used in the example is in the file schooling.txt.

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The baseline model is also implied by the code on the lecture slides on instrumental variables (p. 8–9); it was

```
log(wage) = schooling + exper + exper^2 + black + smsa + south + iqscore,
```

where wage stands for wage (WAGE76 in data), schooling for education (ED76 in data), exper for work experience (EXP76 in data), black for being black (BLACK in data), smsa for living in a metropolitan SMSA (SMSA76 in data) and south for living in the south (SOUTH76 in data). Further details on the model and the data appear on the textbook, chapter 5.4.

(a) It was argued that *schooling*, and thus also *exper* and its square that depend on it, are endogenous because of a missing general ability variable in the model. The file **schooling.txt** contains the variable *iqscore* (an intelligence test score) that can be used as a measure of general ability. Add *iqscore* to the model for the log wage discussed in the video, and estimate the augmented model by OLS. Interpret the coefficient estimates of *schooling* and *iqscore*. What might explain the difference in the the estimate of the coefficient of *schooling* compared to the model excluding *iqscore*?

References/tips: the empirical example in the lectures is covered on slides "Instrumental variables estimator", p. 8–9. In the book, see chapter 5.4. For example code, see schooling.R.

Here, you might think about important unobserved variables that are correlated with both schooling and IQ scores; see also slides 3–5 on "Inconsistent OLS estimator".

The full dataset and its description are available from David Card's homepage; the original dataset includes rows where log wage is missing, has more variables, and some of the variables have slightly different names. The working paper on this topic is freely available.

The dataset contains missing values for the variable *iqscore*. One way to restrict the dataset to rows with no missing values in R is given below.

```
schooling <- read.table('schooling.txt', header = T)
schooling <- subset(schooling, !is.na(iqscore))</pre>
```

(b) Test for heteroskedasticity of the error term using the Breusch-Pagan test in the model in (a). Compute the heteroskedasticity consistent standard errors, and test the significance of the coefficient of schooling. Is the conclusion different compared to the significance test based on assuming homoskedasticity?

References/tips: See the previous week's homework. Note that functions like coeftest also report standard errors when given a covariance matrix; you do not need to report the matrix itself separately.

(c) As pointed out by Verbeek, omitting the ability variable is not the only potential reason for the endogeneity of *schooling*. Estimate the augmented model using *age* (AGE76 in data) and its square and *nearcollege* (NEARC4 in data) as instruments for *exper*, its square and *schooling*. Interpret the coefficient estimates of *schooling* and *igscore*, and compare them to those obtained by OLS in (a).

References/tips: as with example code schooling.R (week 5), you may find ivreg (from the package with the same name) helpful. In particular, consider this bit from ?ivreg:

"Regressors and instruments for ivreg are most easily specified in a formula with two parts on the right-hand side, e.g.,  $y \sim x1 + x2 \mid z1 + z2 + z3$ , where x1 and x2 are the explanatory variables and z1, z2, and z3 are the instrumental variables. Note that exogenous regressors have to be included as instruments for themselves.

For example, if there is one exogenous regressor ex and one endogenous regressor en with instrument in, the appropriate formula would be  $y \sim en + ex \mid in + ex$ . Alternatively, a formula with three parts on the right-hand side can also be used:  $y \sim ex \mid en \mid in$ . The latter is typically more convenient, if there is a large number of exogenous regressors."

See also chapters 5.3.1, 5.3.4, 5.4 in the book, and the slides on instrumental variables estimator.

(d) Interpret the Durbin-Wu-Hausman test result in (c). Is there any evidence of a weak instruments problem?

References/tips: You may find the summary on an object returned by ivreg helpful (use the argument diagnostics=TRUE). For weak instruments and the DWH test, see slide 7 on "Instrumental variables estimator" (week 4) and slides 10–11 on "Generalised instrumental variables estimator" (week 5). In the book, see chapter 5.6.4 and pp. 153–154 in chapter 5.3.1.

(e) Suppose also *iqscore* is endogenous, and estimate the augmented model using as instruments *momseducation* (mother's education, MOMED in data), *dadseducation* (father's education, DADED in data) and *kww* (the score of a general knowledge test, KWW in data) in addition to *age* and its square and *nearcollege*. Interpret the coefficient estimates of *schooling* and *iqscore*, and compare them to those obtained by OLS in (a).

References/tips: see (c).

(f) Interpret the results of the over-identifying restrictions test and Durbin-Wu-Hausman test in (e).

References/tips: again, you may find the summary on an object returned by ivreg helpful. Regarding the over-identifying restrictions test, see slides 8–9 on "Generalised instrumental variables estimator" (week 5), and chapter 5.6.3 in the book.

2. Consider, again, the model for log wage discussed in the video lecture. Estimate the IV model in two stages involving OLS regressions. Verify that the coefficient estimates are identical to those obtained in one stage, but the standard errors are different.

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References/tips: see how a regression for schooling on the left hand side is done in the example code schooling.R. Could you do the same for the other endogenous variables in the model? Note also that if results\_modelX holds your OLS results, you can extract fitted values with fitted(results\_modelX) in R.

The idea of the procedure is outlined on slide 7 (compare to slide 6 as needed) of "Generalized instrumental variables estimator". See also chapter 5.6.2 in the book.

3. The file money.xlsx contains the easonally adjusted quarterly time series of the logarithm of the real money supply,  $m_t$ , real GDP,  $y_t$ , and the 3-month Treasury Bill rate,  $r_t$ , for Canada for the period 1967:3 to 1998:4. In the file, the variables are named as m, y, and r, respectively, while m1 and m2 are  $m_{t-1}$  and  $m_{t-2}$ , respectively, and r1 and r2 are  $r_{t-1}$  and  $r_{t-2}$ , respectively. A conventional money demand function is

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$$m_t = \beta_1 + \beta_2 r_t + \beta_3 y_t + \beta_4 m_{t-1} + \beta_5 m_{t-2} + \varepsilon_t.$$

- (a) Estimate the model by OLS. Interpret the estimation result.
- (b) Test for conditional homoskedasticity of the errors by the Breusch-Pagan test and for first-order error autocorrelation by the Breusch-Godfrey test. Compute the heteroskedaticity consistent covariance matrix estimator, and test for the significance of the coefficients using heteroskedasticity robust standard errors.

References/tips: see the previous week's exercises.

(c) Estimate the model by 2SLS, treating  $r_t$  as an endogenous variable and using  $r_{t-1}$  and  $r_{t-2}$  as additional instruments. Are the estimates much different from the OLS estimates?

References/tips: see 1c).

(d) Show that the model estimated in (c) is over-identified. Test the over-identifying restrictions. Interpret the results of the Durbin-Wu-Hausman test. Is there evidence of a weak instruments problem?

References/tips: see 1d) and 1f).

$$y_i = \beta_1 + \beta_2 x_i + \varepsilon_i$$

Assume that  $x_i$  is independently normally distributed with mean zero and variance  $\sigma_X^2 = 1$ , and the error term  $\varepsilon_i = \rho x_i + \eta_i$ , where  $\eta_i$  is independently normally distributed with mean zero and variance  $\sigma_n^2 = 1$ .

(a) Write down the moment conditions for OLS estimation of  $\beta_1$  and  $\beta_2$ . Show that if  $\rho \neq 0$ , the error term is correlated with the regressor  $x_i$ , and one of the moment conditions is violated.

References/tips: see slide 3 on "Instrumental variables estimator" for moment conditions. For derivation of the correlation, recall that if random variables z and q are independent, then  $\mathbb{E}(zq) = \mathbb{E}q\mathbb{E}z$  (or check 'Some useful results" from problem set 1). Also note that to derive a non-zero correlation, you only need a non-zero covariance, and that covariance can be written as  $\text{Cov}(x_i, \varepsilon_i) = \mathbb{E}x_i\varepsilon_i - \mathbb{E}x_i\mathbb{E}\varepsilon_i$  (see, for example, appendix B.4 in the book).

(b) Let  $\beta_1 = 0.0$  and  $\beta_2 = 1.0$ . Consider three values of  $\rho$ ,  $\rho = 0$ ,  $\rho = 0.25$ , and  $\rho = 0.50$ . For each  $\rho$ , generate S = 5000 samples of size N = 25 from the regression model, and for each generated sample, compute the OLS estimate of  $\beta_2$  and the t-test statistic for  $H_0: \beta_2 = 1.0$  against  $H_1: \beta_2 \neq 1.0$ .

Because 1.0 is the true value of  $\beta_2$ ,  $H_0$  should be rejected in 5% of the replications in the t-test conducted at the 5% level of significance (the nominal size of the test). Compute the rejection rate of the test, i.e., find the proportion of the replications where the absolute value of the t-test statistic exceeds the critical value (1.96).

Plot also the histogram of the OLS estimator of  $\beta_2$ . How does the rejection rate and the distribution of the estimator vary with  $\rho$ ? [Hint: If the model is estimated using the lm() function in R and the result is stored in ols1, the OLS estimate of  $\beta_2$  is obtained as coef(ols1)[2] and the covariance matrix estimator of the OLS estimator as vcov(ols1).]

References/tips: see the previous week's suggested solutions on how to perform t-tests on each iteration of repeatedly simulating data. Standard errors for the regressor x can be extracted by, for example, sqrt(vcov(model)['x','x']) or coef(summary(model))['x','Std. Error']).

If the null hypothesis specifies the coefficient as  $\beta_k^0$ , the coefficient estimate is  $b_k$  and the standard error is  $se(b_k)$ , the t-test statistic is defined as  $\frac{b_k - \beta_k^0}{se(b_k)}$ ; see chapter 2.5.1 in the book or slides 2–3 in week 1's slides Testing hypotheses under the normality assumption.

A histogram can be plotted using the function hist.

(c) Repeat (b) for sample size N = 100, and compare.